

INTRODUCTION

- Count time series data are found in multiple applications. Such time series counts come with inflation and in a bivariate form that captures not only serial dependence within each time series but also interdependence between the two series.
- We constructed a class of bivariate integer-valued time series model [1] by applying either the bivariate Gaussian copula or the bivariate t copula functions.
- This method allows for flexible dependence structures between the two time series and within each one.
- First order Markov chains are considered with zero-inflated Poisson (ZIP), zero-inflated negative binomial (ZINB) and zero-inflated Conway-Maxwell-Poisson (ZICMP) marginals.

SIMULATION RESULTS

Gaussian copula and Student's t copula were selected as candidate copula families with true parameters ($\lambda_1=3$, $\lambda_2=5$, $\omega_1=0.3$, $\omega_2=0.4$, $\delta_1=0.6$, $\delta_2=0.4$, $\rho=0.5$).

| Sample Size | Parameters | Estimate | SE | MSE | MAE |
|-------------|-----------------|----------|-------|-------|-------|
| 100 | $\lambda_1(3)$ | 3.402 | 0.389 | 0.312 | 0.460 |
| | $\omega_1(0.3)$ | 0.333 | 0.083 | 0.008 | 0.070 |
| | $\lambda_2(5)$ | 5.199 | 0.383 | 0.186 | 0.334 |
| | $\omega_2(0.4)$ | 0.403 | 0.069 | 0.005 | 0.054 |
| | $\delta_1(0.6)$ | 0.542 | 0.084 | 0.010 | 0.079 |
| | $\delta_2(0.4)$ | 0.363 | 0.096 | 0.011 | 0.081 |
| | $\rho(0.5)$ | 0.482 | 0.091 | 0.009 | 0.075 |
| | | | | | |
| 300 | $\lambda_1(3)$ | 3.405 | 0.197 | 0.203 | 0.408 |
| | $\omega_1(0.3)$ | 0.338 | 0.045 | 0.003 | 0.047 |
| | $\lambda_2(5)$ | 5.182 | 0.210 | 0.077 | 0.223 |
| | $\omega_2(0.4)$ | 0.406 | 0.039 | 0.002 | 0.031 |
| | $\delta_1(0.6)$ | 0.554 | 0.043 | 0.004 | 0.052 |
| | $\delta_2(0.4)$ | 0.367 | 0.054 | 0.004 | 0.049 |
| | $\rho(0.5)$ | 0.471 | 0.049 | 0.003 | 0.044 |
| | | | | | |
| 500 | $\lambda_1(3)$ | 3.410 | 0.172 | 0.198 | 0.411 |
| | $\omega_1(0.3)$ | 0.341 | 0.037 | 0.003 | 0.046 |
| | $\lambda_2(5)$ | 5.184 | 0.162 | 0.060 | 0.203 |
| | $\omega_2(0.4)$ | 0.408 | 0.029 | 0.001 | 0.025 |
| | $\delta_1(0.6)$ | 0.556 | 0.032 | 0.003 | 0.047 |
| | $\delta_2(0.4)$ | 0.370 | 0.043 | 0.003 | 0.041 |
| | $\rho(0.5)$ | 0.472 | 0.039 | 0.002 | 0.038 |
| | | | | | |

| Sample Size | Parameters | Estimate | St_Dev | MSE | MAE |
|-------------|-----------------|----------|--------|-------|-------|
| 100 | $\lambda_1(3)$ | 2.634 | 0.396 | 0.464 | 0.550 |
| | $\omega_1(0.3)$ | 0.338 | 0.061 | 0.008 | 0.071 |
| | $\kappa_1(0.4)$ | 0.339 | 0.051 | 0.008 | 0.071 |
| | $\lambda_2(5)$ | 4.768 | 1.004 | 3.021 | 1.289 |
| | $\omega_2(0.4)$ | 0.412 | 0.055 | 0.007 | 0.065 |
| | $\kappa_2(0.6)$ | 0.577 | 0.072 | 0.015 | 0.010 |
| | $\delta_1(0.6)$ | 0.576 | 0.071 | 0.016 | 0.101 |
| | $\delta_2(0.4)$ | 0.379 | 0.091 | 0.026 | 0.128 |
| 300 | $\rho(0.5)$ | 0.488 | 0.086 | 0.019 | 0.104 |
| | $\lambda_1(3)$ | 2.813 | 0.234 | 0.234 | 0.437 |
| | $\omega_1(0.3)$ | 0.319 | 0.038 | 0.003 | 0.046 |
| | $\kappa_1(0.4)$ | 0.364 | 0.031 | 0.005 | 0.006 |
| | $\lambda_2(5)$ | 4.825 | 0.687 | 0.501 | 0.559 |
| | $\omega_2(0.4)$ | 0.413 | 0.032 | 0.001 | 0.028 |
| | $\kappa_2(0.6)$ | 0.577 | 0.050 | 0.003 | 0.044 |
| | $\delta_1(0.6)$ | 0.590 | 0.042 | 0.002 | 0.034 |
| 500 | $\delta_2(0.4)$ | 0.387 | 0.052 | 0.003 | 0.042 |
| | $\rho(0.5)$ | 0.493 | 0.052 | 0.003 | 0.042 |
| | $\lambda_1(3)$ | 2.816 | 0.209 | 0.224 | 0.432 |
| | $\omega_1(0.3)$ | 0.318 | 0.031 | 0.002 | 0.039 |
| | $\kappa_1(0.4)$ | 0.368 | 0.028 | 0.005 | 0.067 |
| | $\lambda_2(5)$ | 4.814 | 0.056 | 0.363 | 0.492 |
| | $\omega_2(0.4)$ | 0.411 | 0.028 | 0.001 | 0.024 |
| | $\kappa_2(0.6)$ | 0.574 | 0.042 | 0.002 | 0.040 |
| | $\delta_1(0.6)$ | 0.587 | 0.036 | 0.001 | 0.031 |
| | $\delta_2(0.4)$ | 0.387 | 0.041 | 0.002 | 0.034 |
| | $\rho(0.5)$ | 0.495 | 0.037 | 0.001 | 0.029 |
| | | | | | |

PROPOSED BIVARIATE MODEL

- Consider the following series of 2 – dimensional vector, $\{\mathbf{Y}_t\}_{t=1}^n$, where $\mathbf{Y}_t = (Y_{1t}, Y_{2t})'$ for $t = 1, 2, \dots, n$.
- We assume that each series $\{Y_{1t}\}_{t=1}^n$ and $\{Y_{2t}\}_{t=1}^n$ follows a copula-based first order Markov process.
- The joint probability distribution of Y_{1t} and Y_{2t} given $Y_{1,t-1}$ and $Y_{2,t-1}$, respectively, for $t = 1, \dots, n$, is given by :

$$f(y_{1t}, y_{2t} | y_{1,t-1}, y_{2,t-1}) = \int_{V^{-1}(F_{1,t}^-)}^{V^{-1}(F_{1,t}^+)} \int_{V^{-1}(F_{2,t}^-)}^{V^{-1}(F_{2,t}^+)} V_2(z_1, z_2, R) dz_2 dz_1$$

- V^{-1} denotes either the inverse cdf of the normal distribution or the t-distribution with $V_2(\cdot, R)$ being the bivariate normal or t-distribution, respectively.
- R denotes the correlation matrix associated with the bivariate distribution capturing the cross-sectional dependence between the two series.
- The limits of the integral function are calculated using the the transition cdf of Y_{it} given $Y_{i,t-1}$, for $i = 1, 2$.
- The likelihood function for the proposed model is given by:

$$L(\boldsymbol{\vartheta}, \mathbf{y}) = f(y_{11}, y_{21}) \cdot \prod_{t=2}^n f(y_{1t}, y_{2t} | y_{1,t-1}, y_{2,t-1}),$$

where $\boldsymbol{\vartheta} = (\boldsymbol{\theta}', \delta_1, \delta_2, \rho)'$, is the marginal parameter vector, δ_1 and δ_2 are the copula parameters to deal with the first and second time series, respectively, and finally ρ the bivariate dependence parameter.

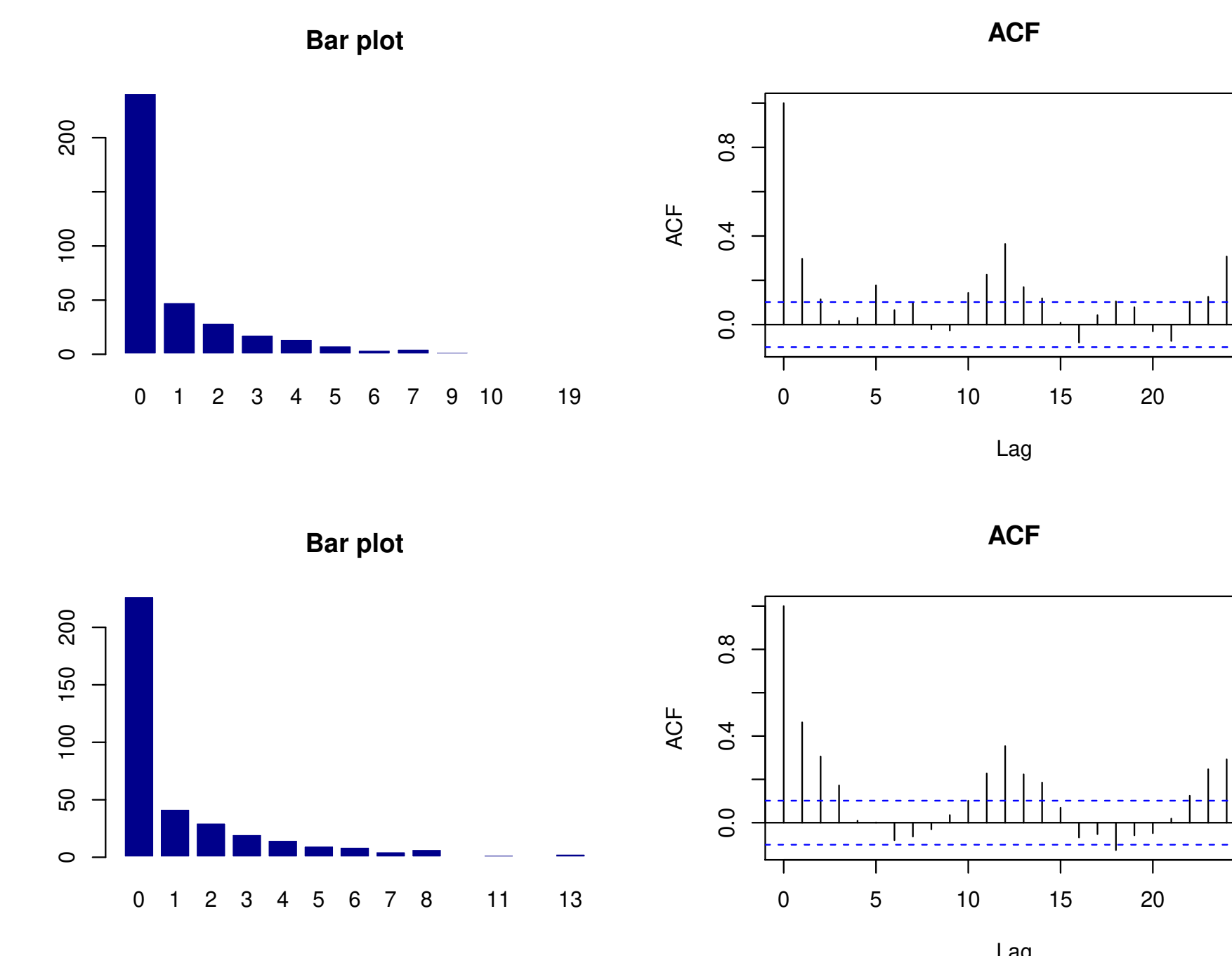
- When calculating ML estimates, the bivariate integral function was evaluated using the standard randomized importance sampling method.
- Taking the inverse of the Fisher information matrix, yields standard errors of the ML estimates.

REFERENCES

- [1] Alqawba M., Fernando D., and Diawara N. A class of copula-based bivariate poisson time series models with applications. *Computation*, 108(9), 2021.
- [2] Eike Christian Brechmann and Claudia Czado. Multivariate time series modeling using the copula autoregressive model. *Applied Stochastic Models in Business and Industry*, 31(4):495–514, 2015.

APPLICATION

The data consists of the monthly counts of strong sandstorms recorded by the AQI airport stations in Eastern Province, Saudi Arabia.



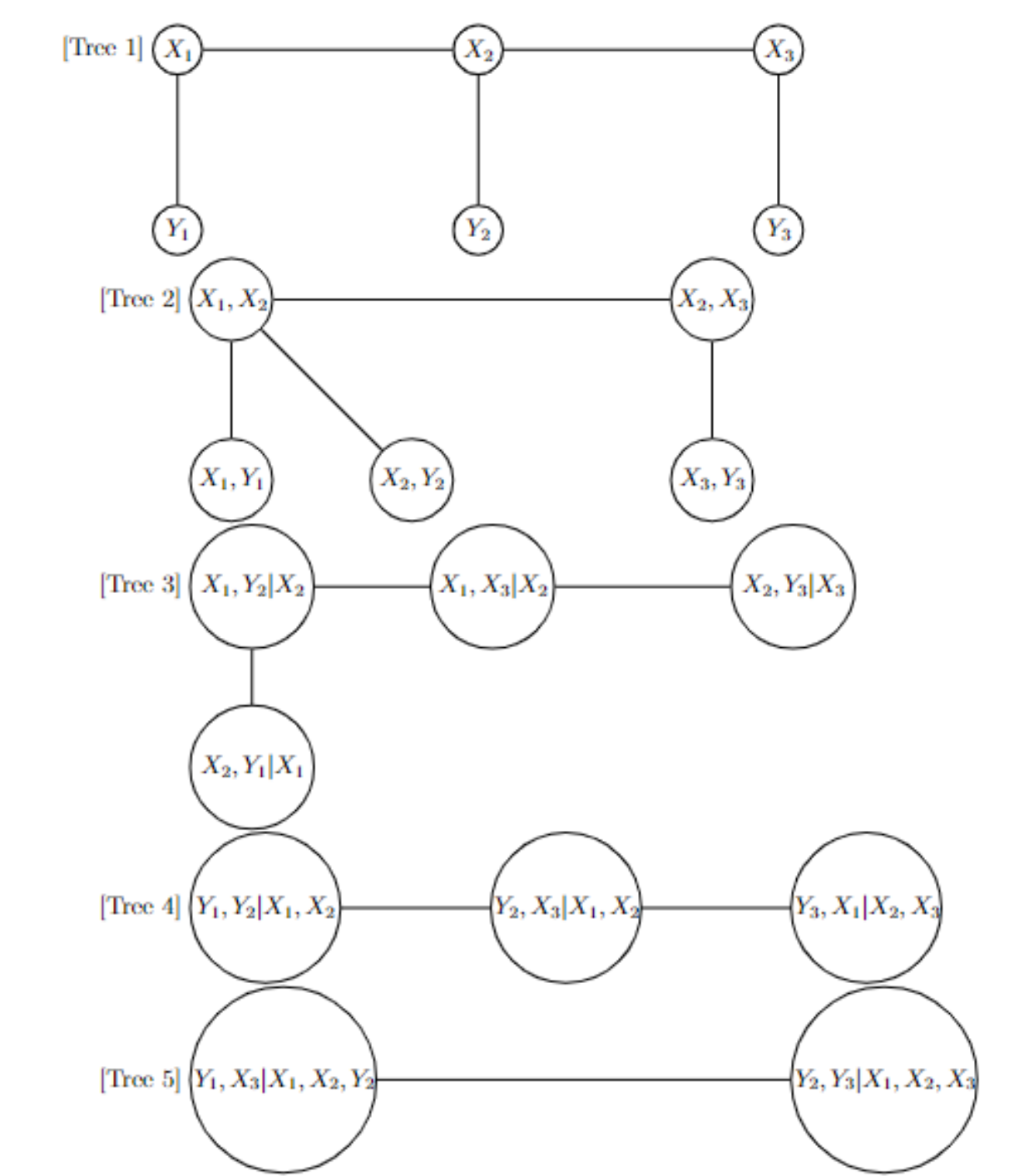
Parameter estimates for the BZINB model.

| Parameter | Estimate | SE |
|-------------|----------|-------|
| λ_1 | 1.758 | 0.197 |
| ω_1 | 0.446 | 0.397 |
| κ_1 | 0.961 | 0.419 |
| λ_2 | 1.946 | 0.295 |
| ω_2 | 0.348 | 0.727 |
| κ_2 | 0.708 | 0.401 |
| δ_1 | 0.412 | 0.059 |
| δ_2 | 0.536 | 0.051 |
| ρ | 0.460 | 0.059 |

THE COPAR MODEL

- The idea is to extend the work of (Brechmann Czado, 2015) [2] to model zero-inflated count time series data.
- They introduced a copula-based model for stationary time series with a Markovian structure which captures both serial dependence and cross sectional dependence in multivariate time series.

Tree structure of 2 dimensional ($\{X_t\}$ and $\{Y_t\}$) COPAR model with 3 time points.



R-vine structure matrix (Morales, 2010) corresponding for the 2 dimensional COPAR model with 3 time points.

$$\begin{pmatrix} Y_3 & & & & \\ Y_1 & X_3 & & & \\ Y_2 & Y_1 & Y_2 & & \\ X_1 & Y_2 & Y_1 & X_2 & \\ X_2 & X_1 & X_1 & Y_1 & Y_1 \\ X_3 & X_2 & X_2 & X_1 & X_1 & X_1 \end{pmatrix}$$

Matrix of copulas for the 2 dimensional COPAR model with 3 time points.

$$\begin{pmatrix} * & & & & \\ C_2^Y & * & & & \\ C_1^Y & C_2^{YX} & * & & \\ C_2^{XY} & C_1^{YX} & C_1^{XY} & * & \\ C_1^{XY} & C_2^X & C_1^{XY} & C_1^{YX} & * \\ C_0^{XY} & C_1^X & C_0^{XY} & C_1^X & C_0^{XY} & * \end{pmatrix}$$

COPAR model of order k (COPAR(k)) is defined, where all pair copulas corresponding to a lag length greater than k are independence copulas.

$$C_{t-s}^X = C_{t-s}^{XY} = C_{t-s}^{YX} = C_{t-s}^Y = \Pi \quad \text{for } t - s > k,$$

where Π denotes the independence copula, respectively.